Portfolios that Contain Risky Assets 15: Cautious Objectives for Markowitz Portfolios

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Portfolios that Contain Risky Assets
Part II: Stochastic Models

Cautious Objectives for Markowitz Portfolios

Introduction

The Kelly criterion says that investors whose objective is to maximize the value of their portfolio over an extended period should maximize its growth rate mean. More precisely, it suggests that such investors should select an allocation $\mathbf{f}$ that maximizes the estimator $\hat{\gamma}(\mathbf{f})$. This suggestion rests upon:

- the validity of an IID model,
- the Law of Large Numbers,
- the estimator $\hat{\gamma}(\mathbf{f})$ is accurate.

If these assumptions hold then the Kelly criterion would be suitable for many young investors, but not for those older investors who depend upon their portfolios for their income. Of course, the last assumption is foolhardy, so even young investors should be more cautious.

Introduction

The Kelly criterion exposes investors to potential downside events from which it might be hard to recover. Older investors who depend upon their portfolios for their income might be drawing down on their portfolio at a rate of 4% per year, hoping that this income stream will last at least 20 years. But if the value of their portfolio is reduced by 40% in a market downturn then their future income stream will be similarly reduced. This puts them in a tough spot if their reduced income no longer covers their fixed expenses. They might feel forced to draw down at 6.5% per year, which would rapidly erode the value of their portfolio.

Here we will develop objective functions that are better suited for more cautious investors. We will do so within the framework of IID models.

Cautious Objectives

Given a set of assets with a return history $\{\mathbf{r}(d)\}_{d=1}^{D}$ and a choice of positive weights $\{w_d\}_{d=1}^{D}$ that sum to one, the Kelly Criterion selects the portfolio allocation $\mathbf{f}$ that maximizes $\hat{\gamma}(\mathbf{f})$ over a class Π of Markowitz portfolio allocations $\mathbf{f}$, where the objective $\hat{\gamma}(\mathbf{f})$ is the sample mean estimator of $\gamma(\mathbf{f})$ given by

$$\hat{\gamma}(\mathbf{f}) = \sum_{d=1}^{D} w_d \log\bigl(1 + r(d,\mathbf{f})\bigr) , \tag{2.1}$$

with

$$r(d,\mathbf{f}) = \mu_{\mathrm{rf}}(\mathbf{f})(1 - \mathbf{1}^{\mathrm{T}}\mathbf{f}) + \mathbf{r}(d)^{\mathrm{T}}\mathbf{f} ,$$
$$\mu_{\mathrm{rf}}(\mathbf{f}) = \begin{cases} \mu_{\mathrm{si}} & \text{for } \mathbf{1}^{\mathrm{T}}\mathbf{f} \le 1 , \\ \mu_{\mathrm{cl}} & \text{for } \mathbf{1}^{\mathrm{T}}\mathbf{f} > 1 . \end{cases} \tag{2.2}$$

Cautious Objectives

Here we present the family of *cautious objectives* that has the form

$$\widehat{\Gamma}^{\chi}(\mathbf{f}) = \hat{\gamma}(\mathbf{f}) - \chi \sqrt{\hat{\theta}(\mathbf{f})}\,, \tag{2.3}$$

where the nonnegative parameter χ is the so-called *caution coefficient* and $\hat{\theta}(\mathbf{f})$ is the sample variance estimator of $\theta(\mathbf{f})$ given by

$$\hat{\theta}(\mathbf{f}) = \frac{1}{1 - \bar{w}_D} \sum_{d=1}^{D} w_d \Big(\log\big(1 + r(d, \mathbf{f})\big) - \hat{\gamma}(\mathbf{f}) \Big)^2 \,, \tag{2.4}$$

with $\bar{w}_D$ defined by

$$\bar{w}_D = \sum_{d=1}^{D} w_d^2 \,. \tag{2.5}$$

Cautious Objectives

Because $\hat{\gamma}(\mathbf{f})$ is a strictly concave function of $\mathbf{f}$, it is clear that $\widehat{\Gamma}^{\chi}(\mathbf{f})$ given by (2.3) will be is a strictly concave function of $\mathbf{f}$ over any bounded set provided that the caution coefficient χ is small enough.

We will soon see that the function $\sqrt{\hat{\theta}(\mathbf{f})}$ is convex over relevant sets of $\mathbf{f}$. In that case for every $\chi > 0$ the additional term in $\widehat{\Gamma}^{\chi}(\mathbf{f})$ will enhance the strict convexity of $\hat{\gamma}(\mathbf{f})$, which helps guard against overbetting.

The choice of a value for the caution coefficient χ is up to each individual investor. It characterizes how cautious the investor wishes to be. Caution can arise from many sources, each of which has to be quantified in order to guide the choice of χ.

Cautious Objectives

We will consider contributions to the caution coefficient from two sources:

- the uncertainty in the sample mean estimator $\hat{\gamma}(\mathbf{f})$ given by (2.1);
- the desire to reduce the impact of downside market events.

The first will be analyzed using the Chebyshev inequality bounds that we developed earlier.

The second will require more information than the Law of Large Numbers provides. However, this additional information can be estimated with the aid of the *Central Limit Theorem.*

Other potential contributions to the caution coefficient will be explored in the projects. These might include our confidence in the IID model or our assesment of economic factors.

Sample Mean Estimator Uncertainty

Recall that the uncertainty in the sample mean estimator $\hat{\gamma}(\mathbf{f})$ can be quantified by the Chebyshev inequality, which shows for every $\delta > \sqrt{\bar{w}_D}$ that

$$\Pr\left\{ |\hat{\gamma}(\mathbf{f}) - \gamma(\mathbf{f})| > \delta \sqrt{\theta(\mathbf{f})} \right\} \leq \frac{\bar{w}_D}{\delta^2} .$$

This can be recast as

$$\Pr\left\{ |\hat{\gamma}(\mathbf{f}) - \gamma(\mathbf{f})| \leq \delta \sqrt{\theta(\mathbf{f})} \right\} \geq 1 - \frac{\bar{w}_D}{\delta^2} . \tag{3.6}$$

This implies that

$$\Pr\left\{ \gamma(\mathbf{f}) \geq \hat{\gamma}(\mathbf{f}) - \delta \sqrt{\theta(\mathbf{f})} \right\} \geq 1 - \frac{\bar{w}_D}{\delta^2} . \tag{3.7}$$

Sample Mean Estimator Uncertainty

This suggests that if $\bar{w}_D/\delta^2$ is small then with high probability

$$\gamma(\mathbf{f}) \geq \hat{\gamma}(\mathbf{f}) - \delta \sqrt{\hat{\theta}(\mathbf{f})} \,. \tag{3.8}$$

Notice that here we have replaced the unknown $\theta(\mathbf{f})$ in (3.7) with its estimator $\hat{\theta}(\mathbf{f})$ given by (2.4).

Remark. Because we see from (2.2) that $r(d, \mathbf{0}) = \mu_{\text{si}}$ for every d, we see from (2.1) and (2.4) that

$$\hat{\gamma}(\mathbf{0}) = \log(1 + \mu_{\text{si}}) \,, \qquad \hat{\theta}(\mathbf{0}) = 0 \,.$$

Therefore because $\gamma(\mathbf{0}) = \log(1 + \mu_{\text{si}})$, we see that inequality (3.8) is an equality for $\mathbf{f} = \mathbf{0}$.

Sample Mean Estimator Uncertainty

Now let $\lambda_\mathrm{e} \in (0,1)$ be the probability that we hope inequality (3.8) holds. By setting

$$\lambda_\mathrm{e} = 1 - \frac{\bar{w}_D}{\delta^2}\,,$$

we obtain

$$\delta = \sqrt{\frac{\bar{w}_D}{1-\lambda_\mathrm{e}}}\,.$$

This suggests that if this sample mean estimator uncertainty was our only concern then we could select the caution coefficient

$$\chi = \sqrt{\frac{\bar{w}_D}{1-\lambda_\mathrm{e}}}\,. \tag{3.9}$$

Of course, the addition of other concerns will lead to a higher value for χ.

Central Limit Theorem

Because the Central Limit Theorem is used to analyze our next concern, we now review it. Let $\{X_d\}_{d=1}^{\infty}$ be any sequence of IID random variables drawn from a probability density $p(X)$ with mean γ and variance $\theta > 0$. Let $\{Y_d\}_{d=1}^{\infty}$ be the sequence of random variables defined by

$$Y_d = \frac{1}{d} \sum_{d'=1}^{d} X_{d'} \qquad \text{for every } d = 1, \cdots, \infty\,. \tag{4.10}$$

Recall that

$$\mathrm{Ex}(Y_d) = \gamma\,, \qquad \mathrm{Var}(Y_d) = \frac{\theta}{d}\,.$$

The Law of Large Numbers says that Y_d is will approach γ as $d \to \infty$. However, it does not say much about how the Y_d are distibuted around γ for fixed d. The Central Limit Theorem gives such information.

Central Limit Theorem

Let $\{Z_d\}_{d=1}^{\infty}$ be the sequence of random variables defined by

$$Z_d = \frac{Y_d - \gamma}{\sqrt{\theta/d}} \qquad \text{for every } d = 1, \cdots, \infty \,.$$

These random variables have been standardized so that

$$\mathrm{Ex}(Z_d) = 0 \,, \qquad \mathrm{Var}(Z_d) = 1 \,.$$

The Central Limit Theorem says that as $d \to \infty$ the limiting distribution of Z_d will be the mean-zero, variance-one normal distribution.

Central Limit Theorem

More precisely, it says that for every $\zeta \in \mathbb{R}$ we have

$$\lim_{d\to\infty} \mathrm{Pr}\{Z_d \geq -\zeta\} = \mathrm{N}(\zeta)\,, \tag{4.11}$$

where $\mathrm{N}(\zeta)$ is the normal cummulative distribution function defined by

$$\mathrm{N}(\zeta) \equiv \int_{-\infty}^{\zeta} \frac{1}{\sqrt{2\pi}}\, e^{-\frac{1}{2}Z^2}\, \mathrm{d}Z = \int_{-\zeta}^{\infty} \frac{1}{\sqrt{2\pi}}\, e^{-\frac{1}{2}Z^2}\, \mathrm{d}Z\,. \tag{4.12}$$

We can express the limit (4.11) in terms of Y_d as

$$\lim_{d\to\infty} \mathrm{Pr}\left\{ Y_d \geq \gamma - \zeta\sqrt{\theta/d} \right\} = \mathrm{N}(\zeta)\,. \tag{4.13}$$

Central Limit Theorem

Remark. The normal cummulative distribution function N is an increasing, continous function that maps $\mathbb{R}$ onto $(0,1)$. It thereby has an inverse N^{-1} that is an increasing, continuous function that maps $(0,1)$ onto $\mathbb{R}$. Both of these functions are infinitely differentiable.

Remark. The power of the Central Limit Theorem is that it assumes so little about the underlying probability density $p(X)$. Specifically, it assumes that

$$\int_{-\infty}^{\infty} X^2 p(X)\,\mathrm{d}X < \infty\,,$$

and that

$$0 < \theta = \int_{-\infty}^{\infty} (X-\gamma)^2 p(X)\,\mathrm{d}X\,, \quad \text{where} \quad \gamma = \int_{-\infty}^{\infty} X\,p(X)\,\mathrm{d}X\,.$$

Central Limit Theorem

The Central Limit Theorem does not estimate how fast the limit (4.13) is approached. Such estimates require additional assumptions about the underlying probability density $p(X)$. The Berry-Esseen Theorem is the simplest such theorem, but it is not covered in most undergraduate probability courses.

The Berry-Esseen Theorem says that there exists $C_{\mathrm{BE}} \in \mathbb{R}$ such that if

$$\rho = \int_{-\infty}^{\infty} |X - \gamma|^3 p(X) \, \mathrm{d}X < \infty \,,$$

then for every $\zeta \in \mathbb{R}$ we have

$$\left| \Pr\left\{ Y_d \geq \gamma - \zeta\sqrt{\theta/d} \right\} - \mathrm{N}(\zeta) \right| \leq C_{\mathrm{BE}} \, \frac{\rho}{\sqrt{d\,\theta^3}} \,. \qquad (4.14)$$

Central Limit Theorem

Remark. The Berry-Esseen Theorem shows that the rate of convergence of the limit (4.13) in the Central Limit Theorem is $d^{-\frac{1}{2}}$ as $d \to \infty$.

Remark. The quantity ρ is the absolute centered third moment of the probability density $p(X)$. The Hölder Inequality shows that it is bounded below by the variance θ of $p(X)$ as

$$\rho \geq \sqrt{\theta^3}\,.$$

Because the ratio $\rho/\sqrt{\theta^3}$ is higher for densities with fatter tails, the error bound (4.14) is bigger for them too.

Remark. The constant C_{BE} is universal because it does not depend upon the probability density $p(X)$. Its best value is known to lie within the interval $(0.4, 0.5)$. Bounding this value is the subject of current research.

Downside Risk Management

The IID model for the Markowitz portfolio with allocation $\mathbf{f}$ has a growth rate mean $\gamma(\mathbf{f})$ and a growth rate variance $\theta(\mathbf{f})$ that are estimated from the return history $\{\mathbf{r}(d)\}_{d=1}^{D}$ by $\hat{\gamma}(\mathbf{f})$ and $\hat{\theta}(\mathbf{f})$ given by (2.1) and (2.4).

Let $\{X_d\}_{d=1}^{\infty}$ be an IID growth rate history drawn from this model and let $\{Y_d\}_{d=1}^{\infty}$ be defined by (4.10). The Law of Large Numbers says that as $d \to \infty$ the values of Y_d become strongly peak around $\gamma(\mathbf{f})$. This behavior seems to be consistent with the idea that a reasonable approach towards portfolio management is to select $\mathbf{f}$ to maximize the estimator $\hat{\gamma}(\mathbf{f})$. However, by taking $\zeta = 0$ in (4.13) we see that the Central Limit Theorem implies

$$\lim_{d\to\infty} \mathrm{Pr}\{Y_d \geq \gamma(\mathbf{f})\} = \mathrm{N}(0) = \tfrac{1}{2}\,.$$

This shows that in the long run the growth rate of a portfolio will exceed $\gamma(\mathbf{f})$ with a probability of only $\frac{1}{2}$. Cautious investors might want the portfolio to exceed the optimized growth rate with a higher probability.

Downside Risk Management

The Central Limit Theorem says that if T is large enough then we can use the approximation

$$\mathrm{Pr}\left\{ Y_T \geq \gamma(\mathbf{f}) - \zeta\sqrt{\theta(\mathbf{f})/T} \right\} \approx \mathrm{N}(\zeta)\,. \tag{5.15}$$

Let $\lambda_{\mathrm{d}} \in (\frac{1}{2}, 1)$ be the probability that we do not want to experience a downside event. Set

$$\zeta = \mathrm{N}^{-1}(\lambda_{\mathrm{d}})\,.$$

Then approximation (5.15) becomes

$$\mathrm{Pr}\left\{ Y_T \geq \gamma(\mathbf{f}) - \frac{\zeta}{\sqrt{T}}\sqrt{\theta} \right\} \approx \lambda_{\mathrm{d}}\,. \tag{5.16}$$

Downside Risk Management

This suggests that if downside tail events were our only concern then we could pick the caution coefficient

$$\chi = \frac{\mathrm{N}^{-1}(\lambda_{\mathrm{d}})}{\sqrt{T}} \,. \tag{5.17}$$

$$\hat{\gamma}(\mathbf{f}) - \frac{\mathrm{N}^{-1}(\lambda)}{\sqrt{T}} \sqrt{\hat{\theta}(\mathbf{f})} \,. \tag{5.18}$$

Downside Risk Management

Remark. Because $\hat{\gamma}(\mathbf{f})$ is a strictly concave function of $\mathbf{f}$, it is clear that $\hat{\Gamma}(\lambda, T, \mathbf{f})$ will be is a strictly concave function of $\mathbf{f}$ over any bounded set provided that $\mathrm{N}^{-1}(\lambda)/\sqrt{T}$ is small enough.

Remark. *The only new assumption we have made in order to construct this objective is that T is large enough for the Central Limit Theorem to yield a good approximation of the distribution of growth rates.*

Remark. Investors often choose T to be the interval at which the portfolio will be rebalanced, regardless of whether T is large enough for the Central Limit Theorem approximation to be valid. If an investor plans to rebalance once a year then $T = 252$, twice a year then $T = 126$, four times a year then $T = 63$, and twelve times a year then $T = 21$. The smaller T, the less likely it is that the Central Limit Theorem approximation is valid.

Downside Risk Management

The idea now will be to select the admissible Markowitz allocation $\mathbf{f}$ *that maximizes* $\hat{\Gamma}(\lambda, T, \mathbf{f})$ *given a choice of* λ *and* T *by the investor. In other words, the objective will be to maximize the growth rate that will be exceeded by the portfolio with probability* λ *when it is held for* T *trading days.* Because $1 - \lambda$ is the fraction of times the investor is willing to experience a downside tail event, the choice of λ reflects how cautious the investor feels. More cautious investors will select a higher λ.

Remark. The caution of an investor can increase with age. Retirees whose portfolio provide an income that covers much of their living expenses will often be extremely cautious. Investors within ten years of retirement may be fairly cautious because they have less time for their portfolio to recover from an economic downturn. In constrast, young investors can be less cautious because they have more time to experience economic upturns and because they are typically far from their peak earning capacity.

Downside Risk Management

Remark. The caution of an investor should also depend on a careful reading of economic factors or an analysis of the historical data. For example, if the historical data shows evidence of a bubble then any investor should be more cautious.

An investor can simply select $\zeta = \mathrm{N}^{-1}(\lambda)$ such that λ is a probability that reflects his or her caution. For example, an investor can select ζ based on the following tabulations

$$\begin{aligned}
&\mathrm{N}\left(\tfrac{1}{4}\right) \approx .5987\,, \quad \mathrm{N}\left(\tfrac{1}{2}\right) \approx .6915\,, \quad \mathrm{N}\left(\tfrac{3}{4}\right) \approx .7734\,, \quad \mathrm{N}(1) \approx .8413\,,\\
&\mathrm{N}\left(\tfrac{5}{4}\right) \approx .8944\,, \quad \mathrm{N}\left(\tfrac{3}{2}\right) \approx .9332\,, \quad \mathrm{N}\left(\tfrac{7}{4}\right) \approx .9599\,, \quad \mathrm{N}(2) \approx .9772\,,\\
&\mathrm{N}\left(\tfrac{9}{4}\right) \approx .9878\,, \quad \mathrm{N}\left(\tfrac{5}{2}\right) \approx .9938\,, \quad \mathrm{N}\left(\tfrac{11}{4}\right) \approx .9970\,, \quad \mathrm{N}(3) \approx .9987\,.
\end{aligned}$$

Downside Risk Management

An investor who is willing to experience a downside tail event roughly

$$
\begin{aligned}
\text{once every two years might select } \zeta &= 0\,,\\
\text{twice every five years might select } \zeta &= \tfrac{1}{4}\,,\\
\text{thrice every ten years might select } \zeta &= \tfrac{1}{2}\,,\\
\text{twice every nine years might select } \zeta &= \tfrac{3}{4}\,,\\
\text{once every six years might select } \zeta &= 1\,,\\
\text{once every ten years might select } \zeta &= \tfrac{5}{4}\,,\\
\text{once every fifteen years might select } \zeta &= \tfrac{3}{2}\,,\\
\text{once every twenty five years might select } \zeta &= \tfrac{7}{4}\,,\\
\text{once every forty four years might select } \zeta &= 2\,.
\end{aligned}
$$

Downside Risk Management

Remark. *We should pick a larger value of ζ whenever our analysis of the historical data gives us less confidence either in the health of the economy, in the calibration of* $\mathbf{m}$ *and* $\mathbf{V}$*, or in the validity of an IID model.* These are the questions that are addressed in the projects.

Remark. This approach is similar to something in financal management called *value at risk*. The finance problem is much harder because the time horizon T considered there is much shorter, typically on the order of days. In that setting the Central Limit Theorem approximation is certainly invalid.

Mean-Variance Estimators of Cautious Objectives

We now use the estimator $\hat{\Gamma}(\lambda, T, \mathbf{f})$ to derive new estimators of $\Gamma(\lambda, T, \mathbf{f})$ in terms of sample estimators of the return mean and variance given by

$$\hat{\mu}(\mathbf{f}) = \mu_{\mathrm{rf}}(\mathbf{f})\big(1 - \mathbf{1}^{\mathrm{T}}\mathbf{f}\big) + \mathbf{m}^{\mathrm{T}}\mathbf{f}\,, \qquad \mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}\,, \tag{6.19}$$

where $\mathbf{m}$ and $\mathbf{V}$ are given by

$$\begin{aligned} \mathbf{m} &= \sum_{d=1}^{D} w_d\, \mathbf{r}(d)\,, \\ \mathbf{V} &= \sum_{d=1}^{D} w_d\, \big(\mathbf{r}(d) - \mathbf{m}\big)\big(\mathbf{r}(d) - \mathbf{m}\big)^{\mathrm{T}}\,. \end{aligned} \tag{6.20}$$

These new mean-variance estimators of $\Gamma(\lambda, T, \mathbf{f})$ will allow us to work within the framework of Markowitz portfolio theory.

Mean-Variance Estimators of Cautious Objectives

Recall that $\hat{\mu}(\mathbf{f})$ is the sample mean of of the history $\{r(d,\mathbf{f})\}_{d=1}^{D}$ and that

$$r(d,\mathbf{f}) - \hat{\mu}(\mathbf{f}) = \tilde{\mathbf{r}}(d)^{\mathrm{T}}\mathbf{f}\,,$$

where $\tilde{\mathbf{r}}(d) = \mathbf{r}(d) - \mathbf{m}$. In words, $\tilde{\mathbf{r}}(d)$ is the deviation of $\mathbf{r}(d)$ from its sample mean $\mathbf{m}$. Then we can write

$$\begin{aligned} x(d,\mathbf{f}) &= \log\big(1 + r(d,\mathbf{f})\big) \\ &= \log\big(1 + \hat{\mu}(\mathbf{f})\big) + \log\left(1 + \frac{\tilde{\mathbf{r}}(d)^{\mathrm{T}}\mathbf{f}}{1 + \hat{\mu}(\mathbf{f})}\right) . \end{aligned} \tag{6.21}$$

We used this expression in the last lecture to derive several estimators of $\hat{\gamma}(\mathbf{f})$. We now use it to derive several estimators of $\hat{\theta}(\mathbf{f})$.

Mean-Variance Estimators of Cautious Objectives

We use the first-order Taylor approximation $\log(1+r) \approx r$ in the second term of (6.21) to obtain

$$x(d,\mathbf{f}) \approx \log\big(1+\hat{\mu}(\mathbf{f})\big) + \frac{\tilde{\mathbf{r}}(d)^{\mathrm{T}}\mathbf{f}}{1+\hat{\mu}(\mathbf{f})}\,. \tag{6.22}$$

This yields the approximation

$$\hat{\theta}(\mathbf{f}) = \sum_{d=1}^{D} \frac{w_d}{1-\bar{w}}\big(x(d,\mathbf{f}) - \hat{\gamma}(\mathbf{f})\big)^2 \approx \frac{1}{1-\bar{w}}\,\frac{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}{(1+\hat{\mu}(\mathbf{f}))^2}\,,$$

which leads to the Taylor estimator

$$\hat{\theta}_{\mathrm{t}}(\mathbf{f}) = \frac{1}{1-\bar{w}}\,\frac{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}{(1+\hat{\mu}(\mathbf{f}))^2}\,. \tag{6.23}$$

Like the Taylor estimator $\hat{\gamma}_{\mathrm{t}}(\mathbf{f})$, this estimator is not well-behaved.

Mean-Variance Estimators of Cautious Objectives

The simplest thing to do is drop the $\hat{\mu}(\mathbf{f})$ term in the denominator of $\hat{\theta}_{\mathrm{t}}(\mathbf{f})$, which leads to the quadratic estimator

$$\hat{\theta}_{\mathrm{q}}(\mathbf{f}) = \frac{1}{1-\bar{w}}\,\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}\,. \tag{6.24}$$

We then introduce the *caution coefficient* χ by

$$\chi = \frac{1}{\sqrt{1-\bar{w}}}\,\frac{\zeta}{\sqrt{T}}\,. \tag{6.25}$$

Typically $\chi < 1$.

Mean-Variance Estimators of Cautious Objectives

When the estimator (6.24) is combined with the parabolic estimator $\hat{\gamma}_{\mathrm{p}}(\mathbf{f})$ we obtain

$$\hat{\Gamma}_{\mathrm{p}}^{\chi}(\mathbf{f}) = \hat{\mu}(\mathbf{f}) - \tfrac{1}{2}\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f} - \chi\sqrt{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}\,. \tag{6.26}$$

When it is combined with the quadratic estimator $\hat{\gamma}_{\mathrm{q}}(\mathbf{f})$ we obtain

$$\begin{gathered} \hat{\Gamma}_{\mathrm{q}}^{\chi}(\mathbf{f}) = \hat{\mu}(\mathbf{f}) - \tfrac{1}{2}\hat{\mu}(\mathbf{f})^2 - \tfrac{1}{2}\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f} - \chi\sqrt{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}\,, \\ \text{over} \quad \hat{\mu}(\mathbf{f}) \leq 1\,. \end{gathered} \tag{6.27}$$

When it is combined with the reasonable estimator $\hat{\gamma}_{\mathrm{r}}(\mathbf{f})$ we obtain

$$\begin{gathered} \hat{\Gamma}_{\mathrm{r}}^{\chi}(\mathbf{f}) = \log\bigl(1 + \hat{\mu}(\mathbf{f})\bigr) - \tfrac{1}{2}\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f} - \chi\sqrt{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}\,, \\ \text{over} \quad 1 + \hat{\mu}(\mathbf{f}) > 0\,. \end{gathered} \tag{6.28}$$

Mean-Variance Estimators of Cautious Objectives

When it is combined with the sensible estimator $\hat{\gamma}_{\mathrm{s}}(\mathbf{f})$ we obtain

$$\hat{\Gamma}_{\mathrm{s}}^{\chi}(\mathbf{f}) = \log\big(1 + \hat{\mu}(\mathbf{f})\big) - \tfrac{1}{2}\,\frac{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}{1 + 2\hat{\mu}(\mathbf{f})} - \chi\,\sqrt{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}\,, \quad \text{over} \quad 1 + 2\hat{\mu}(\mathbf{f}) > 0\,. \tag{6.29}$$

When it is combined with the Taylor estimator $\hat{\gamma}_{\mathrm{t}}(\mathbf{f})$ we obtain

$$\hat{\Gamma}_{\mathrm{t}}^{\chi}(\mathbf{f}) = \log\big(1 + \hat{\mu}(\mathbf{f})\big) - \tfrac{1}{2}\,\frac{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}{(1 + \hat{\mu}(\mathbf{f}))^2} - \chi\,\sqrt{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}\,, \quad \text{over} \quad 1 + \hat{\mu}(\mathbf{f}) \geq \sqrt{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}\,. \tag{6.30}$$

This estimator is strictly concave over its given domain.

Mean-Variance Estimators of Cautious Objectives

Finally, when $\chi < 1$ and the Taylor estimator $\hat{\gamma}_{\mathrm{t}}(\mathbf{f})$ is combined with the Taylor estimator $\hat{\theta}_{\mathrm{t}}(\mathbf{f})$ given by (6.23) to estimate the growth rate exceeded with probability objective, $\hat{\Gamma}(\lambda, T, \mathbf{f})$ given (5.18), then we obtain the ultimate estimator

$$\hat{\Gamma}_{\mathrm{u}}^{\chi}(\mathbf{f}) = \log\big(1 + \hat{\mu}(\mathbf{f})\big) - \frac{1}{2}\frac{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}{(1 + \hat{\mu}(\mathbf{f}))^2} - \chi\,\frac{\sqrt{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}}{1 + \hat{\mu}(\mathbf{f})}\,, \qquad \text{over} \quad 1 + \hat{\mu}(\mathbf{f}) \geq \frac{\sqrt{\mathbf{f}^{\mathrm{T}}\mathbf{V}\mathbf{f}}}{1 - \chi}\,. \tag{6.31}$$

This estimator is strictly concave over its given domain.